

# Association Between Screen Time Exposure in Children at 1 Year of Age and Autism Spectrum Disorder at 3 Years of Age

## The Japan Environment and Children's Study

Megumi Kushima, MA; Reiji Kojima, MD, PhD; Ryoji Shinohara, PhD; Sayaka Horiuchi, MD, DrPH; Sanae Otawa, MA; Tadao Ooka, MD, PhD; Yuka Akiyama, PhD; Kunio Miyake, PhD; Hiroshi Yokomichi, MD, PhD; Zentaro Yamagata, MD, PhD; and the Japan Environment and Children's Study Group

[+ Supplemental content](#)

**IMPORTANCE** It is unclear to what extent the duration of screen time in infancy is associated with the subsequent diagnosis of autism spectrum disorder.

**OBJECTIVE** To examine the association between screen time in infancy and the development of autism spectrum disorder at 3 years of age.

**DESIGN, SETTING, AND PARTICIPANTS** This cohort study analyzed data from mother-child dyads in a large birth cohort in Japan. This study included children born to women recruited between January 2011 and March 2014, and data were analyzed in December 2020. The study was conducted by the Japan Environment and Children's Study Group in collaboration with 15 regional centers across Japan.

**EXPOSURES** Screen time at 1 year of age.

**MAIN OUTCOMES AND MEASURES** The outcome variable, children diagnosed with autism spectrum disorder at 3 years of age, was assessed using a questionnaire administered to mothers of the participating children.

**RESULTS** A total of 84 030 mother-child dyads were analyzed. The prevalence of children with autism spectrum disorder at 3 years of age was 392 per 100 000 (0.4%), and boys were 3 times more likely to have been diagnosed with autism spectrum disorder than were girls. Logistic regression analysis showed that among boys, when "no screen" was the reference, the adjusted odds ratios were as follows: less than 1 hour, odds ratio, 1.38 (95% CI, 0.71-2.69;  $P = .35$ ), 1 hour to less than 2 hours, odds ratio, 2.16 (95% CI, 1.13-4.14;  $P = .02$ ), 2 hours to less than 4 hours, odds ratio, 3.48 (95% CI, 1.83-6.65;  $P < .001$ ), and more than 4 hours, odds ratio, 3.02 (95% CI, 1.44-6.34;  $P = .04$ ). Among girls, however, there was no association between autism spectrum disorder and screen time.

**CONCLUSIONS AND RELEVANCE** Among boys, longer screen time at 1 year of age was significantly associated with autism spectrum disorder at 3 years of age. With the rapid increase in device usage, it is necessary to review the health effects of screen time on infants and to control excessive screen time.

**Author Affiliations:** Center for Birth Cohort Studies, University of Yamanashi, Chuo, Yamanashi, Japan (Kushima, Shinohara, Horiuchi, Otawa, Yamagata); Department of Health Sciences, University of Yamanashi, Chuo, Yamanashi, Japan (Kojima, Ooka, Akiyama, Miyake, Yokomichi, Yamagata).

**Group Information:** The members of the Japan Environment and Children's Study Group appear in Supplement 2.

**Corresponding Author:** Megumi Kushima, MA, Center for Birth Cohort Studies, University of Yamanashi, 1110 Shimokato, Chuo, Yamanashi, 409-3821, Japan ([kumegumi@yamanashi.ac.jp](mailto:kumegumi@yamanashi.ac.jp)).

JAMA Pediatr. doi:10.1001/jamapediatrics.2021.5778  
Published online January 31, 2022.

Autism spectrum disorder (ASD) has been suggested to be associated with congenital factors, such as genomic mutations<sup>1-4</sup> and prenatal, perinatal, and neonatal risk factors.<sup>5-7</sup> In addition, abnormalities in brain morphology and function have been observed in children with ASD since early childhood.<sup>8,9</sup> In studies conducted in 2019 and 2020, it has been reported that as a postnatal environmental factor, duration of screen time may be associated with ASD characteristics<sup>10,11</sup> and brain morphology specific to ASD.<sup>12</sup> Thus, screen time during infancy, a period of rapid development, may be one of the acquired factors that may be associated with ASD.

In 2019, the World Health Organization published guidelines on healthy physical activity, sedentary behavior, and sleep in children younger than 5 years of age, stating that children should not be exposed to screens at 1 year of age or younger.<sup>13</sup> The American Academy of Pediatrics has also recommended that children should not be exposed to screens until they are 18 months of age; warnings about the adverse effects of screen exposure on the health of children have been issued.

In Japan, the most frequent age at diagnosis for ASD is 3.0 years.<sup>14</sup> However, there are few large cohort studies that have focused on prolonged screen exposure and ASD in infancy. Furthermore, amid the recent outbreak of the COVID-19 pandemic, there has been a rapid change in lifestyles, with electronic devices being used as the main channels of communication and social interactions; thus, screen time among children has increased worldwide.<sup>15-17</sup> Amid this social climate, examining the associations of screen exposure with a child's health is an important public health issue.

Therefore, this study aimed to examine the association of screen exposure (an environmental factor) with the development of ASD during early childhood. To achieve this objective, we examined the association between screen time at 1 year of age and the presence or absence of ASD diagnosis at 3 years of age based on parental responses using data derived from a large Japanese birth cohort study (the Japan Environment and Children's Study).

## Methods

### Study Design and Participants

We conducted a large birth cohort study in Japan. The Japan Environment and Children's Study Group operated in collaboration with 15 regional centers across Japan. Its protocol was reviewed and approved by the Ministry of the Environment's Institutional Review Board on Epidemiological Studies and the ethics committees of all the participating institutions. Approximately 100 000 pregnant women were recruited to participate in the study, and all of the participants provided written informed consent.<sup>18</sup> The research was conducted in accordance with the Ethical Guidelines for Medical and Health Research Involving Human Subjects established by the Ministry of Education, Culture, Sports, Science and Technology and the Ministry of Health, Labour

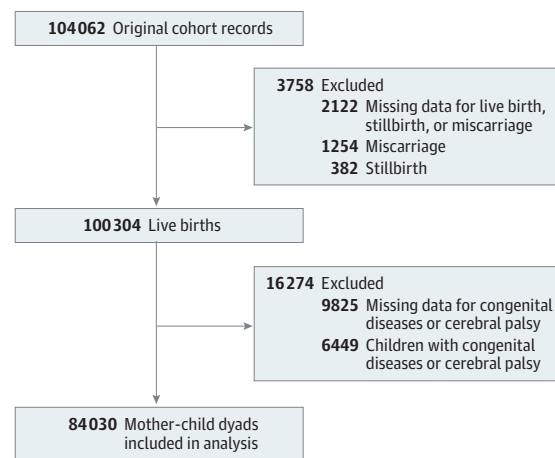
### Key Points

**Question** Is screen-time duration in children at 1 year of age associated with autism spectrum disorder at 3 years of age?

**Findings** A total of 84 030 mother-child dyads were analyzed using data derived from a large birth cohort study conducted in Japan. Among boys, but not girls, longer screen time at 1 year of age was significantly associated with autism spectrum disorder diagnosis at 3 years of age.

**Meaning** Guidance on appropriate screen time in infancy is recommended.

Figure. Selection Process for Participants



and Welfare. The recruitment period was from January 2011 to March 2014, and the data were analyzed during December 2020. The data used were derived from the strictly controlled jecs-ta-20190930-qsn data set, which was released in October 2019. The target population was selected as follows: first, the study included 100 304 live births of the 104 062 fetal records. Consequently, 382 stillbirths and 1254 miscarriages were excluded. We also excluded 2122 individuals with missing data once we calculated live births, stillbirths, and miscarriages. Next, we excluded 6449 children with cerebral palsy, a congenital condition, at 1 year of age, which may have influenced screen time. We also excluded 9825 children with missing data once we calculated congenital diseases or cerebral palsy. Finally, 84 030 mother-child dyads were included in the analysis (Figure).

### Variables

The main exposure variable was screen time at 1 year of age, which was assessed using a questionnaire. When their child turned 1 year of age, the mothers were asked about the number of hours spent per day they let the child watch TV or DVDs. The responses were collected as variables and categorized as, "none (no screen time)," "less than 1 hour," "1 hour or more but less than 2 hours," "2 hours or more but less than 4 hours," and "4 hours or more." When the participating child turned 3 years of age, we asked mothers the

same question. The outcome variable, ASD at 3 years of age, was assessed using a questionnaire. Specifically, mothers were asked the following question when their child turned 3 years of age: "Have they ever been diagnosed with autism spectrum disorder (eg, autism, pervasive developmental disorder, Asperger's syndrome) by a doctor from the age of 2 years until now?" An option of 2 responses was provided, namely, with and without ASD, which were labeled as "yes (ASD)" or "no (no ASD)," respectively.

Previous studies have indicated that maternal nurturing attitudes and abuse may be associated with screen time.<sup>19-21</sup> Therefore, the following factors served as adjustment variables: scores on the Kessler Psychological Distress Scale (K6) and Bonding Scale when the child was 1 year of age; depression, anxiety disorders, integration disorders, and other mental and neurological illnesses; mother's age at delivery; and household income. In addition, when examining the association between ASD at 3 years of age and screen time at 1 year of age, we considered that predisposition to ASD may affect the outcome and may be attributable to reverse causality. We used participant scores on each of the 5 Ages and Stages Questionnaire [R] (ASQ-3) items, which served as adjustment variables to screen for ASD (communication, gross motor skills, fine motor skills, problem-solving, and personal-social scores) at 1 year of age. The ASQ-3 is valid and reliable in screening for developmental delays in children from 1 month of age to 5 1/2 of age, with age-appropriate questions. Having any of these 5 items is below the cutoff value, the line that prompts a visit to a specialist. Although this screening may not fully reflect the characteristics of ASD, its diagnostic accuracy is more than 80% and accurately detects ASD in most cases.<sup>22,23</sup> Because ASD at 1 year of age is currently difficult to diagnose, the ASQ-3 was used in this study, and all of its 5 items were used as adjustment variables.

Each adjustment variable is described below. The Japanese version of the K6 scale was used to assess depressive tendencies,<sup>24</sup> and the cutoff value was set at 5 points or more (Japanese version).<sup>25</sup> A score above the cutoff value of K6 ( $\geq 5$  points) indicated poor mental health. The Japanese version of the Bonding Scale was used to assess mother-child attachment. The scale consists of 10 questions and can yield a maximum score of 30 points; higher scores indicate more negative feelings toward one's baby.<sup>26</sup> However, no clear cutoff value was assigned for this scale; therefore, it was treated as a continuous variable. As younger mothers are at risk of abuse, a cutoff value of 19 years or younger at the time of delivery was used. Because poverty is also a risk factor for abuse, the poverty line reported by the Ministry of Health, Labour and Welfare (annual household income  $\leq$  ¥1.27 million [\$11 060.41 USD]) was used as the standard,<sup>27</sup> and the lowest income item ( $<$  ¥2 million [\$17 417.96 USD]) was used as the cutoff value for responses related to annual household income. The calculation of the poverty line reported by the Ministry of Health, Labour and Welfare is based on the standards of the Organization for Economic Cooperation and Development standards. For the 5 ASQ-3 items, we used the cutoff values specified for the Japanese version.<sup>28</sup> Additionally, because there are sex differ-

ences in ASD,<sup>1,29</sup> sex was used as a stratification variable to examine sex differences in the results.

### Statistical Analysis

First, we aggregated the key variables by screen time at 1 year of age and at 3 years of age, and examined their attributes. Thereafter, odds ratios (ORs) and 95% CIs were calculated to examine the association between screen time and ASD. In addition, because sex differences have been reported in the prevalence of ASD, we examined the interaction by sex and then conducted a sex-stratified analysis. In the trend test in the adjusted model, the categorical variable (screen time) was statistically examined as an ordinal variable (continuous variable). Jonckheere-Terpstra tests were conducted to examine the association between screen time at 1 year of age and 3 years of age. The statistical significance level was set at .05 and 2-tailed. SPSS, version 27 (IBM), was used for statistical analysis.

## Results

### Aggregate Proportions of ASD at 3 Years of Age to Screen Time at 1 Year

A total of 84 030 mother-child dyads were analyzed. In 330 (0.4%) of the 84 030 children included in the analysis, ASD had been diagnosed at 3 years of age (Table 1). Of these, 251 were boys (76.0%) and 79 were girls (24.0%) (eTable 4 in Supplement 1). There were 83 237 responses for screen time at 1 year of age and 74 554 responses for screen time at 3 years of age (Table 1). Irrespective of whether the child had ASD at 3 years of age, at the age 1 year, less than 1 hour was the most commonly provided response for daily screen time (Table 1). The proportion of children with ASD increased as screen time increased (Table 1). The attributes of children and mothers by screen time are shown in Table 1 and eTable 3 in Supplement 1.

### Association Between Screen Time at 1 Year and ASD at 3 Years of Age

Logistic regression analysis of the association between screen time at 1 year and ASD at 3 years of age is shown in Table 2. Longer screen time at 1 year of age was associated with the statistically significantly higher odds of ASD at 3 years of age. Furthermore, the longer screen time at 1 year of age was associated with the statistically significantly higher odds of ASD at 3 years of age in boys. The distribution of screen time was similar across the sexes (Table 1). However, among girls, no association between screen time and ASD was found.

### Association Between Screen Time at 3 Years and ASD at 3 Years of Age

The results of the trend test between screen time at 1 year of age and at 3 years of age showed that screen time at 1 year of age was statistically significantly associated with screen time at 3 years of age (Table 3). Namely, as the screen time increased, the proportion of children with ASD at 3 years of age also increased (Table 2). Moreover, logistic regression analy-

Table 1. Characteristics of the Mother-Child Dyads

Variables	No.	ASD at 3 y of age, No. (%)		Scores on the 5 items of the ASQ-3 at 1 y of age, No. (%)					
		ASD	Child's sex, No. (%)	Communication score <4.53	Gross motor score <9.43	Fine motor score <25.47	Problem-solving score <15.37	Personal-social score <4.95	
			Girls	Boys					
Screen time at 1 y of age, h									
No screen time	8541	19 (5.8)	4014 (9.9)	4527 (10.6)	14 (20.3)	535 (13.5)	520 (12.5)	434 (11.7)	134 (15.9)
<1	27 707	73 (22.3)	13 350 (32.9)	14 356 (33.7)	22 (31.9)	1339 (33.9)	1284 (30.8)	1026 (27.7)	254 (30.1)
1-<2	25 027	99 (30.2)	12 408 (30.5)	12 619 (29.6)	9 (13.0)	1147 (29.0)	1153 (27.6)	1089 (29.4)	248 (29.3)
2-<4	16 560	104 (31.7)	8207 (20.2)	8353 (19.6)	15 (21.7)	718 (18.2)	870 (20.8)	824 (22.3)	151 (17.9)
≥4	5402	33 (10.1)	2644 (6.5)	2757 (6.5)	9 (13.0)	211 (5.3)	348 (8.3)	329 (8.9)	58 (6.9)
Total	83 237	328	40 623	42 612	69	3950	4175	3702	845
Screen time at 3 y of age, h									
No screen time	1253	4 (1.2)	613 (1.7)	640 (1.7)	4 (6.5)	76 (2.1)	72 (1.9)	52 (1.5)	16 (2.1)
<1	17 874	58 (17.6)	8957 (24.6)	8917 (23.4)	13 (21.0)	971 (26.6)	839 (22.0)	702 (20.4)	196 (25.1)
1-<2	33 218	131 (39.7)	16 134 (44.3)	17 084 (44.8)	21 (33.9)	1560 (42.7)	1570 (41.3)	1434 (41.8)	330 (42.3)
2-<4	18 829	109 (33.0)	9114 (25.0)	9715 (25.5)	20 (32.3)	886 (24.2)	1079 (28.4)	1023 (29.8)	201 (25.8)
≥4	3380	28 (8.5)	1620 (4.4)	1760 (4.6)	4 (6.5)	162 (4.4)	245 (6.4)	222 (6.5)	37 (4.7)
Total	74 554	330	36 438	38 116	62	3655	3805	3433	780

Abbreviations: ASD, autism spectrum disorder; ASQ-3, Ages and Stages Questionnaire.

sis reported that screen time at 3 years of age was not associated with ASD at 3 years of age (Table 3).

## Discussion

The main finding of this study was that, among boys, a statistically significant association was found between longer screen time at 1 year of age and ASD at 3 years of age, irrespective of potential maternal maltreatment or predisposition to ASD at 1 year of age. In this study, the prevalence of ASD among 3-year-old children was 0.4%, which is slightly lower than the prevalence of ASD among children younger than 5 years in Asia (0.70%).<sup>30,31</sup> However, given that the prevalence of ASD increases with age, this study's finding is comparable with those of previous studies. In addition, the sex ratio of children with ASD in this study is consistent with what has been observed in previous studies conducted in Japan and abroad.<sup>14,32,33</sup> In Japan, parents who are concerned about their child having a developmental disability often visit medical institutions directly to receive a diagnosis. They could also receive a diagnosis after being advised to visit a medical institution for a possible developmental disability by an infant health checkup (conducted at 4 months, 1½ years, and 3 years of age), nursery school, kindergarten, or elementary school. Medical institutions make a diagnosis of ASD based on the *Diagnostic and Statistical Manual of Mental Disorders* (Fifth Edition).

Despite the World Health Organization and the American Academy of Pediatrics recommendations,<sup>13</sup> 90% of the chil-

dren in this study had been exposed to screens at 1 year of age. Few studies on screen time at 1 year of age have been reported in other countries. According to a survey conducted by the Cabinet Office in Japan, 85.7% of children younger than 1 year and 75.7% of 1-year-old children were using mobile phones, and many of them shared mobile phones with their parents.<sup>34</sup> Thus, at 1 year of age, child-rearing environments may be associated with the development of ASD.

Multivariable analysis of the association between ASD and screen time was conducted to account for the influence of maternal maltreatment and children's predisposition on the result. Among boys, irrespective of their predisposition to ASD at 1 year of age and maternal maltreatment factors, a longer screen time at 1 year of age was associated with ASD at 3 years of age.

In this study, we used the ASQ-3, which has a reported ASD diagnostic accuracy of more than 80%, to adjust for the predisposition of ASD at 1 year of age. However, the diagnostic criteria for ASD include, "Hyper- or hypo-reactivity to sensory input or unusual interest in sensory aspects of the environment" and a strong response to visual information such as lights or movement.<sup>35</sup> This screening is limited because it does not fully reflect the characteristics of ASD. Therefore, we cannot deny the possibility of reverse causality. However, even in that case, screen time can be an effective indicator of ASD during early screening. In addition, this study examined the association between screen time at 1 year of age and ASD by excluding children who received a red flag on any of the 5 ASQ-3 items at 1 year of age (eTable 1 in Supplement 1), thereby helping the ar-

**Table 2. Association Between Screen Time at 3 Years of Age and ASD, Stratified by Child's Sex**

Variables	ASD at 3 y of age, ORs (95% CI)								
	Total			Boys			Girls		
	No.	Model 1, crude	Model 2, adjusted	No.	Model 1, crude	Model 2, adjusted	No.	Model 1, Crude	Model 2, adjusted
Screen time at 1 y of age, <sup>a</sup> h	NA	(n = 74 179)	(n = 61 046)	NA	(n = 37 927)	(n = 31 034)	NA	(n = 36 252)	(n = 30 012)
No screen time	8541	1 [Reference]	1 [Reference]	4527	1 [Reference]	1 [Reference]	4014	1 [Reference]	1 [Reference]
<1	27 707	1.19 (0.72-1.97)	1.16 (0.66-2.03)	14 356	1.45 (0.79-2.64)	1.38 (0.71-2.69)	13 350	0.70 (0.27-1.82)	0.76 (0.27-2.18)
1-<2	25 027	1.81 (1.11-2.96)	1.81 (1.05-3.10)	12 619	2.08 (1.15-3.76)	2.16 (1.13-4.14)	12 408	1.37 (0.56-3.33)	1.28 (0.48-3.46)
2-<4	16 560	2.89 (1.77-4.71)	2.87 (1.68-4.91)	8353	3.40 (1.89-6.11)	3.48 (1.83-6.65)	8207	2.07 (0.85-5.04)	1.93 (0.72-5.20)
≥4	5402	2.80 (1.59-4.93)	2.64 (1.42-4.91)	2757	3.24 (1.66-6.35)	3.02 (1.44-6.34)	2644	2.05 (0.71-5.90)	2.15 (0.67-6.89)
P value for trend	NA	NA	<.001	NA	NA	<.001	NA	NA	.01
Screen time at 3 y of age, <sup>a</sup> h	NA	(n = 74 554)	(n = 61 307)	NA	(n = 38 116)	(n = 31 148)	NA	(n = 36 438)	(n = 30 159)
No screen time	1253	1 [Reference]	1 [Reference]	640	1 [Reference]	1 [Reference]	613	1 [Reference]	1 [Reference]
<1	17 874	1.02 (0.37-2.81)	1.22 (0.38-3.93)	8917	1.03 (0.32-3.33)	1.30 (0.31-5.41)	8957	1.03 (0.14-7.79)	1.10 (0.14-8.53)
1-<2	33 218	1.24 (0.46-3.35)	1.38 (0.44-4.38)	17 084	1.29 (0.41-4.07)	1.60 (0.39-6.54)	16 134	1.06 (0.15-7.83)	0.96 (0.13-7.24)
2-<4	18 829	1.82 (0.67-4.94)	1.79 (0.56-5.70)	9715	1.85 (0.58-5.88)	2.04 (0.50-8.36)	9114	1.68 (0.23-12.44)	1.28 (0.17-9.78)
≥4	3380	2.61 (0.91-7.45)	2.40 (0.71-8.09)	1760	2.19 (0.64-7.47)	2.31 (0.52-10.19)	1620	3.80 (0.49-29.76)	2.64 (0.32-21.99)
P value for trend	NA	NA	.001	NA	NA	.005	NA	NA	.12
Interaction test by sex	<.001	NA	NA	NA	NA	NA	NA	NA	NA

Abbreviations: ASD, autism spectrum disorder; ASQ-3, Ages and Stages Questionnaire; NA, not applicable. anxiety disorders, integration disorders, other mental and neurological illnesses, mother's age at delivery, household income, and scores on the 5 items of the ASQ-3 at 1 year of age.

<sup>a</sup> Adjustment for mothers' scores on the Bonding Scale and Kessler Psychological Distress Scale when the child was 1 year of age, depression,

**Table 3. Association Between Screen Time at 1 Year of Age and Screen Time at 3 Years of Age**

Variable	Screen time at 3 y of age, No.					Total	P value for trend <sup>a</sup>
	No screen	<1 h	1-<2 h	2-<4 h	≥4 h		
Screen time at 1 y of age, h							
No screen time	595	3215	2846	916	92	7664	< .001
<1	385	8475	11 789	3835	305	24 789	NA
1-<2	164	4163	11 237	5949	588	22 101	NA
2-<4	75	1558	5792	6019	1130	14 574	NA
≥4	22	312	1281	1923	1227	4765	NA
Total	1241	17 723	32 945	18 642	3342	73 893	NA

Abbreviation: NA, not applicable.

<sup>a</sup> Jonckheere-Terpstra test.

gument that screen time is a risk factor for the development of ASD.

The results of previous studies that have investigated the association between screen time and ASD in cross-sectional studies are not consistent.<sup>32,36-38</sup> In this study, we examined the association between ASD at 3 years of age, at a single point in time, and screen time at the same age. The results showed no association between screen time and ASD at 3 years of age. This may have been due to the small size of the reference group; therefore, we set the reference group as less than 1 hour and conducted an additional analysis (eTable 2 in Supplement 1).

Consequently, we found that there is a statistically significant difference between boys reported to have 2 to less than 4 hours of screen time. However, the results of the estimates were close to the reference values before they were changed. This indicates that the association of screen time at 1 year of age with ASD diagnosis is still greater than at 3 years of age. This may be because the association with environmental factors on brain development varies with age.

In addition to genetic factors, the role of environmental factors has been noted in ASD. Electromagnetic fields have been cited as an environmental factor associated with health and



screen exposure.<sup>39</sup> Experiments using mice have demonstrated that exposure to high-frequency electromagnetic fields affects neurotransmitters<sup>40</sup> and behavior (hyperactivity and memory impairment)<sup>41</sup> in mice during the developmental period. Additionally, several molecular networks as genetic factors have been associated with the development of ASD, and the core of these molecular networks include  $\alpha$ -amino-3-hydroxy-5-methyl-4-isoxazolepropionic acid receptor (AMPA receptor), protein kinase B (AKT), repressor activator protein 1 (RAP1),  $\gamma$ -aminobutyric acid (GABA), extracellular signal-regulated kinases 1/2 (ERK1/2), methyl-CpG-binding protein 2 (MECP2), brain-derived neurotrophic factor (BDNF), activator protein 1 (AP-1), phosphatase and tensin homolog (PTEN), and ras protein/mitogen-activated protein kinase (RAS/MAPK). Previous studies have reported that low-frequency and high-frequency electrical stimulation, microwave irradiation, and light stimulation of AMPA receptors,<sup>42</sup> Rfn2,<sup>43</sup> GABA,<sup>44</sup> MECP2,<sup>45</sup> and BDNF<sup>1,46-48</sup> are associated with autismlike symptoms. In particular, in infancy when neurodevelopment is active, environmental factors such as electrical stimulation through screens and light stimulation from vision may affect neurodevelopment and de novo sequence alterations.

In this study, we examined the interaction by sex and then conducted a sex-differentiated analysis. The results suggested an association between screen time and ASD only in boys, even if boys and girls have similar screen times (eTable 4 in Supplement 1). This result could be due to the higher prevalence of the disorder in boys. Previous studies on the genetic factors that contribute to the development of ASD have not yet explained the male predominance in ASD.<sup>1</sup> However, among the factors listed as the core genes responsible for the development of ASD, sex differences have been reported for brain-derived neurotrophic factor<sup>49,50</sup> and MECP2.<sup>51</sup> Thus, genetic factors may be involved in the observed sex differences in the association between ASD and screen time.

### Strengths and Limitations

The strength of this study is that it provides new insights into the association of screen time with the onset of ASD in early

childhood, which had not yet been examined. Furthermore, the results of this study are reliable because the analysis was conducted using a large birth cohort data set representative of the Japanese population.<sup>52</sup>

A limitation of this study is that ASD (outcome) and screen time (exposure) were assessed based on parental reports, which may have led to reporting bias. Data on the exact amount of time children are gazing at the screen are not available, and it may differ from what parents report. Medical institutions diagnose ASD based on the *Diagnostic and Statistical Manual of Mental Disorders* (Fifth Edition) but only when ASD can be diagnosed with certainty because it is still in a developmental stage when children are 3 years of age. Therefore, mild cases may not be diagnosed as ASD at 3 years of age, when the child is at a developmental stage, and may be observed in subsequent years. As a result, the study may be biased toward children with severe ASD. To account for potential risks, we conducted a multivariable analysis with ASD predisposition (ASQ-3) at 1 year of age. Additionally, external (eg, prenatal, living, and childcare environment) and internal factors (eg, genetic factors, diseases, and disabilities) other than screen time, which were examined as exposure factors in this study, were not adequately considered. Therefore, this study conducted a sensitivity analysis of unmeasured confounding as an additional analysis (the E-value),<sup>53-57</sup> and it was found that the effect of residual confounding remained (eTable 5 in Supplement 1). Further research is needed to examine other factors involved in the association between ASD and screen time, and the combination of factors at a given time of the year associated with increased risk of ASD.

### Conclusions

In this cohort study, even after accounting for predisposition to ASD at 1 year of age and maternal maltreatment factors, longer screen time at 1 year of age was associated with ASD at 3 years of age in boys. With the rapid increase in the use of devices, it is necessary to review its health effects on infants and control excessive screen time.

#### ARTICLE INFORMATION

**Accepted for Publication:** October 27, 2021.

**Published Online:** January 31, 2022.

doi:10.1001/jamapediatrics.2021.5778

**Open Access:** This is an open access article distributed under the terms of the [CC-BY License](#). © 2022 Kushima M et al. *JAMA Pediatrics*.

**Author Contributions:** Ms Kushima had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

**Concept and design:** Kushima, Yamagata.

**Acquisition, analysis, or interpretation of data:**

All authors.

**Drafting of the manuscript:** Kushima.

**Critical revision of the manuscript for important intellectual content:** All authors.

**Statistical analysis:** Kushima, Shinohara.

**Obtained funding:** Yamagata.

**Administrative, technical, or material support:**

Kushima, Shinohara, Horiuchi, Otawa, Ooka, Yokomichi, Yamagata.

**Supervision:** Yamagata.

**Conflict of Interest Disclosures:** None reported.

**Funding/Support:** This study was funded by the Ministry of Environment, Japan.

**Role of the Funder/Sponsor:** The funder had no role in the design and conduct of the study; collection, management, analysis, and interpretation of the data; preparation, review, or approval of the manuscript; and decision to submit the manuscript for publication.

**Group Information:** Members of the Japan Environment and Children's Study Group are listed in Supplement 2.

**Disclaimer:** The findings and conclusions of this article are solely the responsibility of the authors

and do not represent the official views of the above government body.

**Additional Contributions:** We thank all the participants and cooperating institutions for their contribution to the Japan Environment and Children's Study Group.

#### REFERENCES

1. Liu X, Takumi T. Genomic and genetic aspects of autism spectrum disorder. *Biochem Biophys Res Commun*. 2014;452(2):244-253. doi:10.1016/j.bbrc.2014.08.108
2. Matsumura K, Seiriki K, Okada S, et al. Pathogenic POGZ mutation causes impaired cortical development and reversible autism-like phenotypes. *Nat Commun*. 2020;11(1):859. doi:10.1038/s41467-020-14697-z
3. Sacai H, Sakoori K, Konno K, et al. Autism spectrum disorder-like behavior caused by reduced

- excitatory synaptic transmission in pyramidal neurons of mouse prefrontal cortex. *Nat Commun*. 2020;11(1):5140. doi:10.1038/s41467-020-18861-3
4. Hallmayer J, Cleveland S, Torres A, et al. Genetic heritability and shared environmental factors among twin pairs with autism. *Arch Gen Psychiatry*. 2011;68(11):1095-1102. doi:10.1001/archgenpsychiatry.2011.76
  5. Guinchat V, Thorsen P, Laurent C, Cans C, Bodeau N, Cohen D. Pre-, peri- and neonatal risk factors for autism. *Acta Obstet Gynecol Scand*. 2012;91(3):287-300. doi:10.1111/j.1600-0412.2011.01325.x
  6. Hisle-Gorman E, Susi A, Stokes T, Gorman G, Erdie-Lalena C, Nyland CM. Prenatal, perinatal, and neonatal risk factors of autism spectrum disorder. *Pediatr Res*. 2018;84(2):190-198. doi:10.1038/pr.2018.23
  7. Curran EA, Dalman C, Kearney PM, et al. Association between obstetric mode of delivery and autism spectrum disorder: a population-based sibling design study. *JAMA Psychiatry*. 2015;72(9):935-942. doi:10.1001/jamapsychiatry.2015.0846
  8. Hazlett HC, Poe MD, Gerig G, et al. Early brain overgrowth in autism associated with an increase in cortical surface area before age 2 years. *Arch Gen Psychiatry*. 2011;68(5):467-476. doi:10.1001/archgenpsychiatry.2011.39
  9. Wei L, Zhong S, Nie S, Gong G. Aberrant development of the asymmetry between hemispheric brain white matter networks in autism spectrum disorder. *Eur Neuropsychopharmacol*. 2018;28(1):48-62. doi:10.1016/j.euroneuro.2017.11.018
  10. Madigan S, Browne D, Racine N, Mori C, Tough S. Association between screen time and children's performance on a developmental screening test. *JAMA Pediatr*. 2019;173(3):244-250. doi:10.1001/jamapediatrics.2018.5056
  11. Heffler KF, Sienko DM, Subedi K, McCann KA, Bennett DS. Association of early-life social and digital media experiences with development of autism spectrum disorder-like symptoms. *JAMA Pediatr*. 2020;174(7):690-696. doi:10.1001/jamapediatrics.2020.0230
  12. Hutton JS, Dudley J, Horowitz-Kraus T, DeWitt T, Holland SK. Associations between screen-based media use and brain white matter integrity in preschool-aged children. *JAMA Pediatr*. 2020;174(1):e193869. doi:10.1001/jamapediatrics.2019.3869
  13. World Health Organization. Guidelines on physical activity, sedentary behaviour and sleep for children under 5 years of age. Accessed May 30, 2021. <https://apps.who.int/iris/handle/10665/311664>
  14. Kurasawa S, Tateyama K, Iwanaga R, Ohtoshi T, Nakatani K, Yokoi K. The age at diagnosis of autism spectrum disorder in children in Japan. *Int J Pediatr*. 2018;7(2018):5374725. doi:10.1155/2018/5374725
  15. Aguilar-Farias N, Toledo-Vargas M, Miranda-Marquez S, et al. Sociodemographic predictors of changes in physical activity, screen time, and sleep among toddlers and preschoolers in Chile during the COVID-19 pandemic. *Int J Environ Res Public Health*. 2020;18(1):176. doi:10.3390/ijerph18010176
  16. Schmidt SCE, Anedda B, Burchartz A, et al. Physical activity and screen time of children and adolescents before and during the COVID-19 lockdown in Germany: a natural experiment. *Sci Rep*. 2020;10(1):21780. doi:10.1038/s41598-020-78438-4
  17. Ozturk Eyimaya A, Yalçın İrmak A. Relationship between parenting practices and children's screen time during the COVID-19 pandemic in Turkey. *J Pediatr Nurs*. 2021;56:24-29. doi:10.1016/j.pedn.2020.10.002
  18. Kawamoto T, Nitta H, Murata K, et al; Working Group of the Epidemiological Research for Children's Environmental Health. Rationale and study design of the Japan environment and children's study (JECS). *BMC Public Health*. 2014;14:25. doi:10.1186/1471-2458-14-25
  19. Reck C, Hunt A, Fuchs T, et al. Interactive regulation of affect in postpartum depressed mothers and their infants: an overview. *Psychopathology*. 2004;37(6):272-280. doi:10.1159/000081983
  20. Farrell CA, Flegler EW, Monuteaux MC, Wilson CR, Christian CW, Lee LK. Community poverty and child abuse fatalities in the United States. *Pediatrics*. 2017;139(5):e20161616. doi:10.1542/peds.2016-1616
  21. Zhou Y. Understanding why mothers abuse their children: the case of Japan. *J Health Care Soc*. 2019;29(1):119-134. doi:10.4091/iken.2019.001
  22. Hardy S, Haisley L, Manning C, Fein D. Can screening with the ages and stages questionnaire detect autism? *J Dev Behav Pediatr*. 2015;36(7):536-543. doi:10.1097/DBP.0000000000000201
  23. Beacham C, Reid M, Bradshaw J, et al. Screening for autism spectrum disorder: profiles of children who are missed. *J Dev Behav Pediatr*. 2018;39(9):673-682. doi:10.1097/DBP.0000000000000607
  24. Furukawa TA, Kawakami N, Saitoh M, et al. The performance of the Japanese version of the K6 and K10 in the World Mental Health Survey Japan. *Int J Methods Psychiatr Res*. 2008;17(3):152-158. doi:10.1002/mp.257
  25. Sakurai K, Nishi A, Kondo K, Yanagida K, Kawakami N. Screening performance of K6/K10 and other screening instruments for mood and anxiety disorders in Japan. *Psychiatry Clin Neurosci*. 2011;65(5):434-441. doi:10.1111/j.1440-1819.2011.02236.x
  26. Yoshida K, Yamashita H, Conroy S, Marks M, Kumar C. A Japanese version of Mother-to-Infant Bonding Scale: factor structure, longitudinal changes and links with maternal mood during the early postnatal period in Japanese mothers. *Arch Womens Ment Health*. 2012;15(5):343-352. doi:10.1007/s00737-012-0291-1
  27. Ministry of Health, Labour and Welfare. Comprehensive survey of living conditions (2019). Accessed June 1, 2021. <https://www.mhlw.go.jp/toukei/saikin/hw/k-tyosa/k-tyosa19/index.html>
  28. Mezawa H, Aoki S, Nakayama SF, et al. Psychometric profile of the ages and stages questionnaires, Japanese translation. *Pediatr Int*. 2019;61(11):1086-1095. doi:10.1111/ped.13990
  29. Hiller RM, Young RL, Weber N. Sex differences in pre-diagnosis concerns for children later diagnosed with autism spectrum disorder. *Autism*. 2016;20(1):75-84. doi:10.1177/1362361314568899
  30. Olusanya BO, Davis AC, Wertlieb D, et al; Global Research on Developmental Disabilities Collaborators. Developmental disabilities among children younger than 5 years in 195 countries and territories, 1990-2016: a systematic analysis for the Global Burden of Disease Study 2016. *Lancet Glob Health*. 2018;6(10):e1100-e1121. doi:10.1016/S2214-109X(18)30309-7
  31. Saito M, Hirota T, Sakamoto Y, et al. Prevalence and cumulative incidence of autism spectrum disorders and the patterns of co-occurring neurodevelopmental disorders in a total population sample of 5-year-old children. *Mol Autism*. 2020;11(1):35. doi:10.1186/s13229-020-00342-5
  32. Montes G. Children with autism spectrum disorder and screen time: results from a large, nationally representative US study. *Acad Pediatr*. 2016;16(2):122-128. doi:10.1016/j.acap.2015.08.007
  33. Ros-Demarize R, Bradley C, Kanne SM, et al. ASD symptoms in toddlers and preschoolers: an examination of sex differences. *Autism Res*. 2020;13(1):157-166. doi:10.1002/aur.2241
  34. Cabinet Office GoJ. Survey of youth internet usage environment (2019). Accessed June 1, 2021. <https://www8.cao.go.jp/youth/youth-harm/chousa/r01/net-jittai/pdf-index.html>
  35. American Psychiatric Association. *Diagnostic and statistical manual of mental disorders*. 5th ed. American Psychiatric Association Publishing; 2013.
  36. Healy S, Haegele JA, Grenier M, Garcia JM. Physical activity, screen-time behavior, and obesity among 13-year olds in Ireland with and without autism spectrum disorder. *J Autism Dev Disord*. 2017;47(1):49-57. doi:10.1007/s10803-016-2920-4
  37. Must A, Phillips SM, Curtin C, et al. Comparison of sedentary behaviors between children with autism spectrum disorders and typically developing children. *Autism*. 2014;18(4):376-384. doi:10.1177/136236131479039
  38. Healy S, Garcia JM, Haegele JA. Environmental factors associated with physical activity and screen time among children with and without autism spectrum disorder. *J Autism Dev Disord*. 2020;50(5):1572-1579. doi:10.1007/s10803-018-3818-0
  39. International Commission on Non-Ionizing Radiation Protection. ICNIRP statement on the "Guidelines for limiting exposure to time-varying electric, magnetic, and electromagnetic fields (up to 300 GHz)". *Health Phys*. 2009;97(3):257-258. doi:10.1097/HP.0b013e3181aff9db
  40. Aldad TS, Gan G, Gao XB, Taylor HS. Fetal radiofrequency radiation exposure from 800-1900 mhz-rated cellular telephones affects neurodevelopment and behavior in mice. *Sci Rep*. 2012;2:312. doi:10.1038/srep00312
  41. Kim JH, Huh YH, Lee JH, Jung JY, Ahn SC, Kim HR. Early exposure to radiofrequency electromagnetic fields at 1850 MHz affects auditory circuits in early postnatal mice. *Sci Rep*. 2019;9(1):377. doi:10.1038/s41598-018-36868-1
  42. Takemoto K, Iwanari H, Tada H, et al. Optical inactivation of synaptic AMPA receptors erases fear memory. *Nat Biotechnol*. 2017;35(1):38-47. doi:10.1038/nbt.3710
  43. Morimura N, Yasuda H, Yamaguchi K, et al. Autism-like behaviours and enhanced memory formation and synaptic plasticity in Lrnf2/SALM1-deficient mice. *Nat Commun*. 2017;8:15800. doi:10.1038/ncomms15800
  44. Qiao S, Peng R, Yan H, et al. Reduction of phosphorylated synapsin I (ser-553) leads to spatial memory impairment by attenuating GABA release after microwave exposure in Wistar rats. *PLoS One*.

2014;9(4):e95503. doi:10.1371/journal.pone.0095503

45. Hao S, Tang B, Wu Z, et al. Forniceal deep brain stimulation rescues hippocampal memory in Rett syndrome mice. *Nature*. 2015;526(7573):430-434. doi:10.1038/nature15694

46. Kimura T, Kaneko F, Iwamoto E, Saitoh S, Yamada T. Neuromuscular electrical stimulation increases serum brain-derived neurotrophic factor in humans. *Exp Brain Res*. 2019;237(1):47-56. doi:10.1007/s00221-018-5396-y

47. Baek A, Park EJ, Kim SY, et al. High-frequency repetitive magnetic stimulation enhances the expression of brain-derived neurotrophic factor through activation of Ca<sup>2+</sup>-calmodulin-dependent protein kinase II-cAMP-response element-binding protein pathway. *Front Neurol*. 2018;9:285. doi:10.3389/fneur.2018.00285

48. Croll SD, Suri C, Compton DL, et al. Brain-derived neurotrophic factor transgenic mice exhibit passive avoidance deficits, increased seizure severity and in vitro hyperexcitability in the hippocampus and entorhinal cortex. *Neuroscience*.

1999;93(4):1491-1506. doi:10.1016/S0306-4522(99)00296-1

49. McAllister BB, Bihelek N, Mychasiuk R, Dyck RH. Brain-derived neurotrophic factor and TrkB levels in mice that lack vesicular zinc: effects of age and sex. *Neuroscience*. 2020;425:90-100. doi:10.1016/j.neuroscience.2019.11.009

50. Chan CB, Ye K. Sex differences in brain-derived neurotrophic factor signaling and functions. *J Neurosci Res*. 2017;95(1-2):328-335. doi:10.1002/jnr.23863

51. Plummer JT, Evgrafov OV, Bergman MY, et al. Transcriptional regulation of the MET receptor tyrosine kinase gene by MeCP2 and sex-specific expression in autism and Rett syndrome. *Transl Psychiatry*. 2013;3(10):e316. doi:10.1038/tp.2013.91

52. Michikawa T, Nitta H, Nakayama SF, et al; Japan Environment and Children's Study Group. The Japan Environment and Children's Study (JECS): a preliminary report on selected characteristics of approximately 10 000 pregnant women recruited during the first year of the study. *J Epidemiol*. 2015;25(6):452-458. doi:10.2188/jea.JE20140186

53. Haneuse S, VanderWeele TJ, Arterburn D. Using the e-value to assess the potential effect of unmeasured confounding in observational studies. *JAMA*. 2019;321(6):602-603. doi:10.1001/jama.2018.21554

54. Ding P, VanderWeele TJ. Sensitivity analysis without assumptions. *Epidemiology*. 2016;27(3):368-377. doi:10.1097/EDE.0000000000000457

55. VanderWeele TJ, Ding P. Sensitivity analysis in observational research: introducing the E-value. *Ann Intern Med*. 2017;167(4):268-274. doi:10.7326/M16-2607

56. Mathur MB, Ding P, Riddell CA, VanderWeele TJ. Website and R package for computing E-values. *Epidemiology*. 2018;29(5):e45-e47. doi:10.1097/EDE.0000000000000864

57. VanderWeele TJ, Ding P. Sensitivity analysis in observational research: introducing the E-value. *Ann Intern Med*. 2017;167(4):268-274. doi:10.7326/M16-2607